

1 Questions from lecture notes

1. The key statistical assumption that we made to make the supervised inductive batch learning problem well-defined is:

TRAINING & DEPLOYMENT DATA ARE SAMPLES FROM SAME LIKELIHOOD.

[0.5 Mark]

2. The Bayes optimal classifier with 0-1 loss is the MODE of the underlying posterior likelihood (written as a function of input).

[0.5 Mark]

3. The component of the generalization error that is independent of the training set size is called: model error.

[0.5 Mark]

4. The formula for loss in logistic regression is given by: $l(\text{sign}(a), \text{sign}(b)) \equiv \frac{\log(1 + e^{-ab})}{\log 2}$.

[0.5 Mark]

5. Suppose the underlying likelihood in a linear regression problem satisfies the equations:

$$Y = \sum_{i=1}^n \alpha_i X_i + N, \quad \mathbb{E}[NX] = 0, \quad (1)$$

where X, Y are the input, output. The necessary and sufficient conditions for parameters of a Bayes optimal linear regressor to be same as α are: $\mathbb{E}[XX^T] \succ 0$.

[1 Mark]

6. From bias-variance tradeoff discussion in linear models it is clear that $n \rightarrow \infty, m \rightarrow \infty$ is sufficient for good generalization. Here, n, m are the number of parameters and training set size respectively. But should n diverge like \sqrt{m} or like m or like m^2 ? \sqrt{m} . Fill in this blank with one of the three functions \sqrt{m}, m, m^2 .

[0.5 Mark]

7. In applications of quantum information theory, the space (manifold) of positive definite (pd) matrices is often encountered and the standard loss function is the squared-Bures-Wasserstein metric: loss between A and B is given by $[\text{Tr } A + \text{Tr } B - 2 \text{Tr}(A^{1/2} B A^{1/2})^{1/2}]$. While the standard loss on \mathbb{R} is the squared-loss. Consider a regression problem with input space as that of pd matrices and real outputs, employing standard loss. The Bayes optimal, f^* , for this regression problem is defined by:

$$f^*(x) \equiv \underset{y \in \mathbb{R}}{\text{argmin}} \mathbb{E}[(y - \gamma)^2 / x]$$

Fill this blank with an expression involving the expression for the loss.

[1 Mark]

8. Consider a linear regression problem¹ where the underlying likelihood is such that $p(x, y) = p(x)p(y)$. Assume the mean and variance with $p(x)$ are 3,9 respectively. Assume the mean and variance with $p(y)$ are 4,36 respectively. Then, the simplified expression for the Bayes optimal linear regressor is:

$$f_{\mathcal{L}}^*(x) = \underline{2/3} x.$$

Fill this blank with a numeric constant.

[1.5 Marks]

9. Consider a linear regression problem² with training data (input,output pairs): $\mathcal{D} = \{(1, 2), (3, 4)\}$. Then, the simplified expression for the ERM linear regressor is:

$$\hat{f}_{\mathcal{L}}(x) = \underline{1.4} x.$$

Fill this blank with a numeric constant.

[1.5 Marks]

2 Derivations done in lectures

10. Derive a simplified expression for the Bayes optimal restricted to the linear model over inputs with squared loss using only projection theorem as done in lectures.

$\begin{aligned} \omega^* &= \underset{\omega \in \mathbb{R}^n}{\operatorname{argmin}} E\{(\omega^T X - Y)^2\} \\ &= \underset{\omega \in \mathbb{R}^n}{\operatorname{argmin}} \ \omega^T X - Y\ _2^2 \\ &= \underset{\omega \in \mathbb{R}^n}{\operatorname{argmin}} \ \omega^T E\{XX^T\}^{1/2} \tilde{X} - Y\ _2 \\ &= \underset{\omega \in \mathbb{R}^n}{\operatorname{argmin}} \ \tilde{X} = E\{XX^T\}^{-1/2} X\ _2 \end{aligned}$	$\begin{aligned} &= E\{XX^T\}^{-1/2} \underset{\tilde{\omega} \in \mathbb{R}^n}{\operatorname{argmin}} \ \tilde{\omega}^T \tilde{X} - Y\ _2 \\ &= E\{XX^T\}^{-1/2} E\{\tilde{X} Y\} \quad \downarrow \\ &= E\{XX^T\}^{-1} E\{XY\} \quad \text{Proj. theorem.} \end{aligned}$
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[2 Marks]

3 Problems from other course page resources

11. Consider a problem where the underlying likelihood is defined by $p(x|y) \sim \mathcal{N}(\mu_y, \Sigma_y)$, $y = 1, 0$. Let $p(y) = \begin{cases} 0.4 & y = 1 \\ 0.6 & y = 0 \end{cases}$. Assume the loss is the 0-1 loss. Derive a simplified expression for the Bayes optimal.

¹feature map $\phi(x) = x$.

²feature map $\phi(x) = x$.

$$p(\epsilon/n) \propto e^{-\frac{1}{2}(n-\mu_1)^T \Sigma_1^{-1} (n-\mu_1) + \ln 0.6} \quad f^*(n) = 1 \text{ if } \text{sign}(g(n)) \leq 0, 0 \text{ else}$$

$$p(0/n) \propto e^{-\frac{1}{2}(n-\mu_0)^T \Sigma_0^{-1} (n-\mu_0) + \ln 0.6} \quad g(n) = \frac{1}{2} x^T (\Sigma_1^{-1} - \Sigma_0^{-1}) x - x^T (\Sigma_1^{-1} \mu_1 - \Sigma_0^{-1} \mu_0) - \frac{1}{2} \ln \frac{|\Sigma_0|}{|\Sigma_1|}$$

[1 Mark]

If $\Sigma_1 = \Sigma_0$, then show that the model error with linear model over the feature map, $\phi(x) = \begin{bmatrix} x \\ 1 \end{bmatrix}$, is exactly zero.

$$g(x) \text{ becomes } -x^T (\Sigma_1^{-1} \mu_1 - \Sigma_0^{-1} \mu_0) + \ln 0.6 + \frac{1}{2} \mu_1^T \Sigma_1^{-1} \mu_1 - \frac{1}{2} \mu_0^T \Sigma_0^{-1} \mu_0 - \frac{1}{2} \ln \frac{|\Sigma_0|}{|\Sigma_1|}$$

[0.5 Mark]

Prove that this condition is not necessary for the model error being zero³. In other words, provide a different condition on $\mu_1, \mu_0, \Sigma_1, \Sigma_0$, such that the model error is exactly zero.

model error is zero iff quadratic reduces to an affine or linear or constant in terms of its sign.

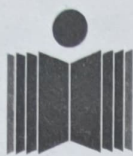
For alternate, $\Sigma_1 \neq \Sigma_0$, i.e. sign of quadratic ~~be~~ a constant

\Leftrightarrow discriminant ≤ 0 ,
 $x^T A x + 2G^T x + C \geq 0 \forall x \Leftrightarrow C - G^T A^{-1} G \geq 0$. So Alternate is (A is invertible or 0)

$\Sigma_1^{-1} - \Sigma_0^{-1} > 0, \mu_1^T \Sigma_1^{-1} \mu_1 - \mu_0^T \Sigma_0^{-1} \mu_0 + 2 \ln 0.6 - \ln \frac{|\Sigma_0|}{|\Sigma_1|}$
 $-(\Sigma_0^{-1} \mu_0 - \Sigma_1^{-1} \mu_0)^T (\Sigma_1^{-1} - \Sigma_0^{-1})^{-1} (\Sigma_0^{-1} \mu_0 - \Sigma_1^{-1} \mu_1) \geq 0$

[4 Marks]

³is the solution provided for the practice problems partially wrong? Go home and think about the necessary and sufficient conditions for 0 model error. If you think you got them, meet me to discuss.



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Roll No. :

① First line on Pg 12 of summary notes

② First line, record para, Pg 16 of notes

③ First line, record para, Pg 20 of notes

④ Second last line, Second last para, Pg 32 of notes

⑤ last para Pg 23 - first para Pg 24 of notes

On

⑥ Second last para in Pg 23 of notes

~~Mainly in notes~~

X_1, \dots, X_n are lin independent $\Rightarrow E[xx^T] \neq 0$

Marks to enter answer.

⑥ Based on Eqn. (10.2) and (11.1) in notes

$$\frac{n}{m} \rightarrow 0 \quad \frac{\sqrt{m}}{m} \rightarrow 0 \quad \frac{m}{m} \rightarrow ? \quad \frac{m^3}{m} \rightarrow \infty$$

⑦ Eqn (6.1) in notes. Substi. tute square loss.

⑧ Eqn (8.7) $w^* = E[xx^T]^{-1} E[xy]$ from notes

$$E[xy] = E[x]E[y] \quad (\because x \perp y)$$

$$= 3 \times 4 = 12$$

$$E\{xx^T\} = E\{x^2\} = \text{var}(x) + (E\{x\})^2$$

$$= 9 + 3^2 = 18$$

$$\therefore \omega^* = \frac{12}{18} = \frac{2}{3} \quad \therefore f_{\omega}^*(x) = \omega^* x$$

$$= \frac{2}{3}x$$

9 From notes (8.8) $\hat{\omega} = \left(\frac{1}{n} \sum_i x_i x_i^T \right)^{-1} \left(\frac{1}{n} \sum_i x_i y_i \right)$

$$= \left(\frac{1}{2} (1 \times 1 + 3 \times 3) \right)^{-1} \left(\frac{1}{2} (1 \times 2 + 3 \times 4) \right)$$

$$= \frac{14}{10} = 1.4$$

$$\therefore \hat{f}_{\omega}(x) = \hat{\omega} x = 1.4x.$$

10 Problem 3 on Pg 3 on hand-written derivations.

11 $p(y|x) \propto p(x|y)p(y)$

$$p(x|y) \propto e^{-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1}(x-\mu_1)} \frac{0.4}{|\Sigma_1|^{1/2}}$$

$$p(y) \propto e^{-\frac{1}{2}(y-\mu_1)^T \Sigma_1^{-1}(y-\mu_1) + \ln 0.4 - \frac{1}{2} \ln |\Sigma_1|}$$

$$= e^{-\frac{1}{2}(x-\mu_0)^T \Sigma_0^{-1}(x-\mu_0) + \ln 0.6 - \frac{1}{2} \ln |\Sigma_0|}$$

11 by $p(x|y) \propto e^{-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1}(x-\mu_1) + \ln 0.4 - \frac{1}{2} \ln |\Sigma_1|}$

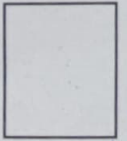
$$p(x|y) \geq p(x|0) \Leftrightarrow -\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1}(x-\mu_1) + \ln 0.4 - \frac{1}{2} \ln |\Sigma_1|$$

$$\geq -\frac{1}{2}(x-\mu_0)^T \Sigma_0^{-1}(x-\mu_0) + \ln 0.6 - \frac{1}{2} \ln |\Sigma_0|$$



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$$\therefore f^*(x) = \begin{cases} 1 & \text{if } \frac{1}{2} x^T (\Sigma_1 - \Sigma_0) x - x^T (\Sigma_1 \mu_1 - \Sigma_0 \mu_0) + \frac{1}{2} \mu_1^T \Sigma_1^{-1} \mu_1 - \frac{1}{2} \mu_0^T \Sigma_0^{-1} \mu_0 + \ln 0.6 \cdot 0.4 \leq 0 \\ 0 & \text{else} \end{cases}$$

$$-\frac{1}{2} \ln \frac{|\Sigma_0|}{|\Sigma_1|}$$

if $\Sigma_1 = \Sigma_0$, then

$$f^*(w) = \begin{cases} 1 & \text{if } -x^T (\Sigma_1^{-1} \mu_1 - \Sigma_0^{-1} \mu_0) + \frac{1}{2} \mu_1^T \Sigma_1^{-1} \mu_1 - \frac{1}{2} \mu_0^T \Sigma_0^{-1} \mu_0 + \ln \frac{0.6}{0.4} - \frac{1}{2} \ln \frac{|\Sigma_0|}{|\Sigma_1|} \leq 0 \\ 0 & \text{else} \end{cases}$$

i.e. $f^*(w) = \text{sign}(-w^T x + b_0^*)$ ($-1 \rightarrow 0$)

\therefore Model error is zero.

Can model ever be zero \rightarrow quadratic \rightarrow linear/affine
 \rightarrow quadratic \rightarrow const.

Can quads of quadratic behave like constant without $\Sigma_1 = \Sigma_0$?

Yes! if its a non-negative quadratic } Discriminant ≤ 0
or non-positive quadratic.

2 marks if you write this.

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$$x^T A x + 2b^T x + c \geq 0 \quad \forall x$$

$$\Leftrightarrow \|A^{1/2}x + A^{-1/2}b\|^2 + c - b^T A^{-1}b \geq 0 \quad \forall x \quad (\text{assuming } A \succ 0)$$

$$\Leftrightarrow c - b^T A^{-1}b \geq 0$$

↓
If you are interested
in knowing what
happens if $A \not\succ 0$
then please meet
me.
as if A is singular
only.

∴ Alternate condition is

$$\Sigma_1^{-1} - \Sigma_0^{-1} \succ 0, \quad \mu_1^T \Sigma_1^{-1} \mu_1 - \mu_0^T \Sigma_0^{-1} \mu_0 + 2 \ln \frac{0.6}{0.4} = \ln \frac{|\Sigma_0|}{|\Sigma_1|}$$

$$- (\Sigma_0^{-1} \mu_0 - \Sigma_1^{-1} \mu_1)^T (\Sigma_1^{-1} - \Sigma_0^{-1}) (\Sigma_0^{-1} \mu_0 - \Sigma_1^{-1} \mu_1) \geq 0$$

or

See Assignment Week 3 for comprehensive answer.